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Inter- and Intra-rater Reliability of a Modified Approach to Scoring the Balance Error Scoring System (BESS) in Undistracted and Distracted Conditions

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Running head: Reliability of modified BESS approach

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Number of words: 3,496

Number of figures: 1

Number of tables: 7

ACKNOWLEDGEMENTS

Thanks to Dr. Eadric Bressel and Dr. Richard Gordin for their assistance with study design and statistical analyses. We would also like to thank Dale Mildenberger, MS, LAT, ATC and Dr. Jonathan Finnoff for developing the modified BESS used in this study.

1 ABSTRACT

2 **Context:** We have developed a modified approach to scoring balance errors with the Balance

3 Error Scoring System (BESS) by eliminating two of the six original error criteria.

4 **Objective:** To measure the inter- and intra-rater reliability of this modified BESS approach.

5 Because of the potential auditory distractions that may be present in a clinical setting, we also

6 measured reliability for distracted rating conditions.

7 **Design:** Cross-sectional.

8 **Setting**: University sports medicine clinic.

9 Patients or Other Participants: Board of Certification (BOC) certified athletic trainers as raters
 10 (n=6).

Intervention: Raters used a modified approach to scoring the BESS to score 32 different sets of pre-recorded videos of balancers completing the six stance positions of the BESS. They first completed this in an undistracted condition, then one week later in a distracted condition as they

14 performed a concurrent auditory vigilance task.

15 **Main Outcome Measures:** Two-way, random, absolute agreement intraclass correlation

16 coefficients (ICC) measured the inter- and intra-rater reliability of the modified approach for

17 undistracted and distracted conditions. Repeated measures ANOVAs tested for effects of

18 condition (undistracted vs. distracted) and also BESS stance position (e.g. single-leg, foam) on

19 the raters' reported number of errors.

20 **Results:** For the undistracted condition, inter- and intra-rater reliability ICCs were 0.93 and

21 0.92, respectively. For the distracted condition, inter- and intra-rater reliability ICCs were 0.89

22 and 0.92, respectively. There was no significant effect of condition (undistracted vs. distracted)

23 on the reported number of errors in the total error score or in the individual stance position

24 scores.

25 Conclusions: This modified approach to scoring the BESS is reliable, even when used in a
 26 potentially distracting environment.

Key Words: balance; reliability; distraction; Balance Error Scoring System

29 INTRODUCTION

30 Reliable balance assessments are necessary in clinical settings to document musculoskeletal injury recovery.¹ determine fall risk.² and decide post-concussion return-to-31 32 play³. Although there are a variety of balance assessment tools, athletic trainers may use the Balance Error Scoring System (BESS).³⁻⁶ The BESS includes three stances (double-leg, single-33 34 leg, and tandem) on two separate surfaces (firm ground and foam), yielding six different stance 35 positions.⁷ This test requires individuals to stand in each position for 20 seconds with their eyes 36 closed and their hands on their hips. To assess an individual's balance using the BESS, a 37 clinician counts the number of errors the individual makes in each stance position. There are six 38 criteria for what constitutes a "balance error" in the original version of the BESS: 1) lifting hands 39 off of hips; 2) opening eyes; 3) stepping, stumbling, or falling out of position; 4) taking >5 40 seconds to return to position; 5) lifting forefoot or heel; or 6) abducting/flexing hip joint >30 41 degrees. The clinician records the number of errors made in each stance position, and then 42 sums each position to yield a total BESS score. The original BESS is valid, can be performed quickly in a small space, requires little equipment, and is inexpensive.⁷ It has been shown, 43 however, to have "poor to moderate" inter- and intra-rater reliability,⁸⁻¹⁰ according to general 44 guidelines.¹¹ One possible explanation for the "poor to moderate"¹¹ reliability of the original 45 BESS is the number of possible error criteria.⁸ 46

47 We have modified the criteria for counting balance errors within the BESS criteria by 48 removing two of the six original error criteria. This modified approach (mBESS) aims to increase 49 the reliability of the original BESS, based on the assumption that clinicians are not highly 50 reliable in visually confirming whether a balancer has lifted his/her forefoot or heel, or has 51 abducted or flexed the hip joint more than 30 degrees. Thus, the mBESS counts only the first 52 four criteria from the BESS as an error, and does not count the last two described above. By 53 minimizing these possible error counts, clinicians using this modified approach may be more 54 reliable at detecting changes in balance performance. Regardless of which approach is used,

55 however, clinicians typically rate the BESS in environments that contain a number of potential 56 auditory distractions, such as a busy athletic training room or a noisy sideline. Distraction can 57 affect balance performance during the BESS.¹² but less is known about the extent to which 58 distraction affects rating ability. If a clinician is required to attend to another task at the same 59 time as rating the BESS, his/her ability to reliably rate balance errors may decline. Thus, the 60 purpose of this study was to measure the inter- and intra-rater reliability of a modified approach 61 to scoring the balance error scoring system (BESS) in undistracted and distracted conditions. 62 We hypothesized that both the inter- and intra-rater reliability of this modified approach would be 63 "good" for the undistracted condition, but only "poor to moderate" for the distracted condition. 64 We also hypothesized that the reported number of errors would be different for the undistracted 65 condition compared to the distracted condition.

66

67 METHODS

68 **Participants**

69 Six Board of Certification (BOC) certified athletic trainers participated in this study 70 (mean±SD age: 29.04±7.87 yrs). Participants used a modified approach to scoring balance 71 errors in the BESS to rate balance performance of 32 (out of 64 total) pre-recorded sets of 72 videos of individuals completing all six stance positions of the BESS, and are therefore referred 73 to as 'raters.' Details regarding these pre-recorded sets of videos and the modified BESS 74 (mBESS) scoring approach are described below in Experimental Design. Mean and SD years of 75 certification was 4.81±6.15 across raters. We first determined the reliability of the mBESS in an 76 undistracted condition (Time 1). We then re-evaluated reliability measures in a distracted 77 condition using an auditory vigilance task one week later (Time 2; description of auditory 78 vigilance task below). All raters provided informed consent prior to participation in compliance 79 with the Institutional Review Board at Utah State University.

81 Experimental design

82 General methods

83 All raters (n=6) sat in front of a computer screen in a guiet room and scored sets of pre-84 recorded balance videos using only four of the six original BESS criteria. Table 1 illustrates the 85 differences in criteria between the original and modified error criteria. If the balancer made 86 multiple errors at the same time, such as opening the eyes and falling out of position, only one 87 error was recorded. Also, once the balancer made an error according to the modified approach, 88 the rater was instructed not to record another error until the balancer had returned to the correct 89 stance position. To ensure that the raters used the correct criteria in their scoring, they were 90 asked to list from memory the criteria that they used to score balance errors for each video at 91 the end of both Time 1 and Time 2. Raters viewed 32 (out of 64) pre-recorded videos within a 92 PowerPoint (Microsoft, Redmond, WA) presentation that had been prepared previously by the 93 experimenter. At both Time 1 and Time 2, raters completed their scoring while wearing 94 headphones, regardless of whether they were subjected to the auditory vigilance task or not. 95 Prior to rater recruitment for this study, we recruited 32 'balancers' who were free from a 96 lower extremity injury in the previous year, and had no history of any physician-diagnosed 97 concussion. Balancers provided informed consent for their participation in this study. Balancers 98 were then video-recorded in the frontal plane as they completed all six stance positions of the 99 BESS: double-leg, firm (ground); single-leg, firm; tandem, firm; double-leg, foam; single-leg, 100 foam; tandem, foam. Each stance position was held for 20 seconds. Additional details regarding these conditions and the test itself have been published previously.^{7,13} All videos were archived 101 102 by balancer and by stance position, such that each balancer yielded six separate 20-sec video 103 files, which equaled one set of videos. For reasons beyond the scope of this study, the 32 104 balancers were video-recorded again a week later in the same fashion, yielding a total of 64 105 sets of videos. A random selection of 32 sets were then inserted offline by the experimenter into 106 PowerPoint presentations for the raters to view, with each 20-sec video separated by a blank

slide to allow the raters time to record the corresponding error score on a paper data collection
sheet. The composition of video sets within the PowerPoint presentations enabled the testing of
inter- and intra-rater reliability of the mBESS approach.

110

111 Time 1

Raters scored 16 sets of videos, were given a five-minute rest period, and then scored another 16 sets of videos. Raters were undistracted during Time 1. After each video, raters were instructed to write down the total number of errors they observed on their data collection sheet. These data were used to determine the inter-rater reliability of the mBESS approach before any potential distractions were introduced in the study.

117

118 Time 2

119 One week later, raters again scored 16 sets of videos, were given a five-minute rest 120 period, and then scored another 16 sets of videos. To determine the effect of distraction on the 121 a) reliability and b) rater performance of the mBESS approach, raters scored one of the two sets 122 of videos (either the first or second set) while performing a concurrent auditory vigilance task 123 that served as a potential distractor to the rater.

124

125 Distracted rating conditions with an auditory vigilance task

We used an auditory vigilance task to potentially distract the raters as they used the mBESS to score balance errors. Auditory vigilance has been shown previously to significantly distract healthy young adults¹⁴ and adults with stroke^{15,16} as they performed a concurrent motor task (i.e. under dual-task conditions), regardless of general intelligence level.^{17,18} The auditory vigilance task in this study required raters to listen to 20-sec recordings of 35-letter sequences. Each 35-letter sequence consisted of a random series of the same four letters (A, G, M, and O). The sequence began as soon as the balancer began the condition, followed by 35 letters (at

133 1.75 Hz), and ran continuously for the duration of the stance position. Prior to each sequence, 134 raters were instructed to pay attention to the number of times a target letter was heard. The 135 target letter was A, G, M, or O, and was changed for each trial. Immediately after each trial, 136 raters were asked to record the number of times a target letter was heard on a paper data 137 collection sheet. Raters were asked to both score the videos to the best of their ability, as well 138 as get the number of target letters correct. All sequences were recorded live using an external 139 microphone (Gigaware Omnidirectional model 33-119, Ignition L.P., Dallas, TX) and played at a 140 comfortable volume through headphones (Sony MDR-V700). All audio files were embedded in 141 the PowerPoint that included the videos. Raters were given two familiarization trials of the 142 auditory vigilance task prior to doing so while simultaneously rating the 16 sets of balance 143 videos at Time 2. During two more trials, to gather a baseline auditory vigilance score, raters 144 remained seated, with eyes closed, and focused only on the number of times a target letter was 145 read.

146

147 Data analysis

Raters reported the number of balance errors using the mBESS approach for each stance position: double-leg, firm (ground); single-leg, firm; tandem, firm; double-leg, foam; single-leg, foam; tandem, foam. The reported numbers of errors from the individual stance positions were also summed to yield a total 'composite' score for each set of videos. These individual stance positions, along with the total composite score, were used to compute reliability measures and to compare rater performance.

154

155 Statistical analysis

156 To measure the inter- and intra-rater reliability of the mBESS approach for both 157 undistracted and distracted conditions, we used two-way, random, absolute agreement 158 intraclass correlation coefficients $(ICC_{2,1})^{11,19}$ All ICC values and confidence intervals were 159 determined using SPSS software version 21 (IBM, Armonk, NY). Inter-rater reliability for 160 undistracted conditions was computed across video sets 1-16 in Time 1 and Time 2. Intra-rater 161 reliability for undistracted conditions was computed within raters across Time 1 and Time 2, 162 using only the error reports from the 16 sets of videos viewed without distraction. Inter-rater 163 reliability for distracted conditions was computed across raters in Time 2, using only the error 164 reports from the 16 sets of videos viewed while simultaneously performing the auditory vigilance 165 task. Intra-rater reliability for distracted conditions was computed within raters using the 16 sets 166 of videos from Time 1 (undistracted) that corresponded with the same video sets rated under 167 distracted conditions from Time 2. All ICC values are reported as 0.00-1.00, with larger values 168 indicating higher reliability. Reliabilities were computed for all six separate BESS stance 169 positions as well as total composite scores.

170 To test whether the reported number of errors was different in the undistracted condition 171 compared to the distracted condition, we compared mean total composite scores across raters 172 using a one-way repeated measures ANOVA with condition (undistracted vs. distracted) as a 173 within-subject factor. Based on our hypothesis, we expected no main effect of condition on the 174 reported number of errors. To test whether the auditory vigilance task affected the raters' 175 scoring of a particular stance position compared to another position, we used a 2x4 repeated 176 measures ANOVA with condition (undistracted vs. distracted) and stance position (single-leg, 177 firm vs. tandem, firm vs. single-leg, foam vs. tandem, foam) as within-subject factors. While there are six separate stance positions in the BESS protocol,⁷ the two double-leg stance 178 179 positions (firm and foam) yielded zero reported errors (see Results). Thus, only the four error-180 inducing stance positions were included in our ANOVA. We expected significant main effects of 181 condition and stance position on the reported number of errors, yet no interaction between 182 condition and stance position. Significant effects were tested post hoc using Tukey-Kramer Honestly Significant Different (HSD) tests.^{20,21} Multiple comparisons were accounted for using 183

- the Kackar-Harville correction.²² All ANOVAs and subsequent posthoc analyses were performed
 with JMP 8.0 (SAS Institute Inc., Carey, NC), with an alpha level of .05.
- 186

187 **RESULTS**

188 What is the reliability of the modified BESS approach for the undistracted condition?

189 Inter-rater reliability

190Table 2 shows the inter-rater intraclass correlation coefficient (ICC) values and191confidence intervals (CI) for each stance position and for the total score when the raters were192undistracted. These values are reported for Time 1 (i.e. initial rating session). The inter-rater193ICC value for the total score was 0.93. The values for both double-leg stance positions (firm and194foam) were zero, as all raters reported no errors. Of the other four stance positions, only single-195leg, foam had an inter-rater ICC value below a "good" range.¹¹

196Table 3 shows the inter-rater ICC values and CIs for each stance position and for the197total score when the raters were undistracted. These values are reported for Time 2 (i.e. second198rating session). The inter-rater ICC value for the total score was also 0.93. The values for both199double-leg stance positions (firm and foam) were zero, as all raters again reported no errors. All200other four stance positions had inter-rater ICC values within the "good" range.¹¹

201

202 Intra-rater reliability

Table 4 shows the intra-rater ICC values and CIs for each stance position and for the total score when the raters were undistracted. These values are reported within rater between Time 1 and Time 2 (i.e. initial and second rating sessions). The intra-rater ICC value for the total score was 0.92. The values for both double-leg stance positions (firm and foam) were zero, as all raters reported no errors. All other four stance positions had intra-rater ICC values within the "good" range.¹¹

210 What is the reliability of the modified BESS approach for the distracted condition?

211 Inter-rater reliability

Table 5 shows the inter-rater ICC values and CIs for each stance position and for the total score when the raters were distracted. These values are reported for Time 2 (i.e. second rating session w/ distractor). The inter-rater ICC value for the total score was 0.89. The values for both double-leg stance positions (firm and foam) were zero, as all raters reported no errors. Of the other four stance positions, only single-leg, foam had an inter-rater ICC value below a "good" range.¹¹

218

219 Intra-rater reliability

Table 6 shows the intra-rater ICC values and CIs for each stance position and for the total score when the raters were distracted. These values are reported for the comparison between Time 1 (undistracted) and Time 2 (distracted) within rater. The intra-rater ICC value for the total score was 0.92. The values for both double-leg stance positions (firm and foam) were zero, as all raters reported no errors. Of the other four stance positions, only single-leg, foam had an inter-rater ICC value below a "good" range.¹¹

226

Was the reported number of errors different between the undistracted and distractedconditions?

Figure 1A shows the reported number of errors for the total BESS score in undistracted and distracted conditions. Our one-way repeated measures ANOVA reported no significant effect of condition (undistracted vs. distracted) on the reported number of errors in the total error score ($F_{1,5} = .045$; p=.83). Given that ICC values varied by BESS stance position (see Tables 2-6), it is plausible that different stance positions could be differentially affected when the rater concurrently performed the auditory vigilance task. Our 2x4 repeated-measures ANOVA reported no significant main effect of condition ($F_{1,5} = .101$; p=.75) nor an interaction between

condition and stance position (F _{3,5}=.057; p=.98) on the reported number of errors. There was,

237 however, a significant main effect of stance position on reported number of errors (F

238 _{3,5}=262.831; p<.0001; Fig. 1B). Post-hoc analyses revealed that the reported number of errors in

all error-inducing stance positions (single leg, firm; tandem, firm; single leg, foam; tandem,

foam) were significantly different from each other (p<.05).

241

242 Were these results due to balancer and rater characteristics?

243 To ensure that the results reported above were not due to the order in which the raters 244 viewed the sets of videos, or the characteristics of the balancers or raters, we repeated the 245 same experiment in three additional raters with a mean±SD age of 26.47±2.65 years and a 246 mean±SD years of experience of 2.74±1.66, and also provided prior informed consent. These 247 additional raters viewed other PowerPoint presentations containing 32 sets of videos that were 248 different from those used in the above results. ICC values and CIs from these additional raters 249 are shown in Table 7A-E. Note the similarities between this table and Tables 2-6, suggesting that the "good"¹¹ reliability of the mBESS in this study was not due to certain balancer 250 251 characteristics.

252

253 **DISCUSSION**

254 The purpose of this study was to measure the inter- and intra-rater reliability of a 255 modified approach to scoring the balance error scoring system (BESS) for undistracted and 256 distracted conditions, as measured by intraclass correlation coefficient (ICC) values. Although we hypothesized "poor to moderate"¹¹ reliability for the distracted condition, we found that the 257 258 modified approach had "good"¹¹ inter- and intra-rater reliability, regardless of condition. General guidelines for interpreting ICC values indicate that values >0.75 as having "good"¹¹ reliability, 259 and those below 0.75 as having "poor to moderate"¹¹ reliability. Moreover, the raters' reported 260 261 number of errors was not significantly different between the undistracted and distracted

262 conditions. Thus, this modified approach to scoring the BESS is reliable, even when used in a
 263 potentially distracting environment.

264 The presence of distractions has been shown to affect balance performance. Numerous studies have reported changes in postural sway,^{23,24} fall risk,²⁵ and postural responses to 265 perturbations²⁶ under dual-task conditions. Even performance on the BESS itself has been 266 267 shown to decline in situations when the balancer is susceptible to distraction.¹² Typically, if a 268 distraction is present enough to disrupt the balancer, it could also disrupt the rater. Given that 269 the reliability and the raters' performance were comparable, however, between undistracted and 270 distracted conditions, our findings suggest that modified approach to rating is a reliable balance 271 assessment regardless of whether a distraction is present or not.

When rating the pre-recorded sets of videos at Time 1 and at Time 2, the raters maintained "good"¹¹ inter-rater ICC values for the total score and for all but one stance position: single-leg, foam at Time 1. The ICC value for this stance position and time point fell within the "poor to moderate"¹¹ range as 0.68, but by Time 2 was categorized as "good"¹¹ with a value of 0.82. This improvement in inter-rater reliability of the single-leg, foam stance position from Time 1 to Time 2 may be a) an effect of learning within raters or b) an effect of stance position difficulty within balancers, as described below.

279 Learning effects have been documented for the original BESS approach, such that balancers improve their total scores the more times they complete the BESS.^{6,27,28} Balancing on 280 281 the foam surface is particularly susceptible to such learning effects. Even in healthy adults, 282 learning to stand on the foam surface requires several consecutive trials that are longer than 20 seconds.²⁹ Our results for the foam surface are likely not attributable to this balancer learning 283 284 effect, however, as all balancers in this study completed the BESS only two times, separated by 285 one week. The change in reliability may have instead been due to rater learning between Time 286 1 and Time 2, yet all raters in this study had two or more years of clinical experience and had 287 used this modified approach in a clinical setting previously. Moreover, if there was a rater

learning effect, then all stance positions should have shown an increase in reliability, rather than
only one. Future studies are needed, however, to identify the minimum number of times a
clinician needs to have administered this modified BESS approach, as well as any other rated
assessment, to reliably use it.

292 Alternatively, the lower ICC values for the single-leg, foam stance position may have 293 been due to a higher number of errors committed by the balancer (and subsequently reported 294 by the rater). Findings from this study are consistent with many others that report significantly 295 higher error rates in the single-leg, foam position compared to all other positions in healthy adults^{7,10,30-33} and concussed individuals.³² This suggests that the balancers who were video-296 297 recorded in this study completed the BESS using similar balance strategies to those in other 298 studies. Regardless, though, if the more 'difficult' stance positions of the BESS (i.e. on foam or 299 in single and tandem stances) influenced the reliability of modified BESS approach, then the 300 ICC values should scale inversely with stance difficulty. The tables of results described earlier 301 do not consistently show this trend, however, with the ICC values often being even higher for 302 the foam surface relative to the firm surface.

Although our results suggest that the modified BESS approach is reliable even when the rater is performing a concurrent task that may be distracting, participants in this study who served as raters were rating pre-recorded videos of healthy adult balancers who were free of musculoskeletal injury and concussion history. Future studies are needed to determine if the "good"¹¹ reliability of this modified BESS approach is preserved when the balancer is injured or concussed.

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396 LEGEND TO FIGURES

- 397 Figure 1. Mean reported number of errors in undistracted and distracted conditions for (A) total
- 398 score and (B) each stance position. Error bars indicate standard error.

Table 1. Criteria for Scoring Errors using the Original and Modified BESS criteria.

Table 5. Distracted Inter-Rater Reliability ICC's and CI's

10	Table 5. Distracted inter-Ro	aler Renability ICC S and Cr S		
	Stance Position	ICC	CI	
	Double Leg Firm	0.00 ^a	0.000 ^a	
	Single Leg Firm	0.86	0.741-0.938	
	Tandem Firm	0.90	0.816-0.958	
	Double Leg Foam	0.00 ^a	0.000 ^a	
	Single Leg Foam	0.56	0.344-0.780	
	Tandem Foam	0.91	0.834-0.965	
	Total Score	0.89	0.769-0.958	

^aBoth double leg stance ICCs and CIs could not be calculated due to a lack of errors.

Table 6. Undistracted vs. Distracted Intra-Rater Reliability ICC's and CI's

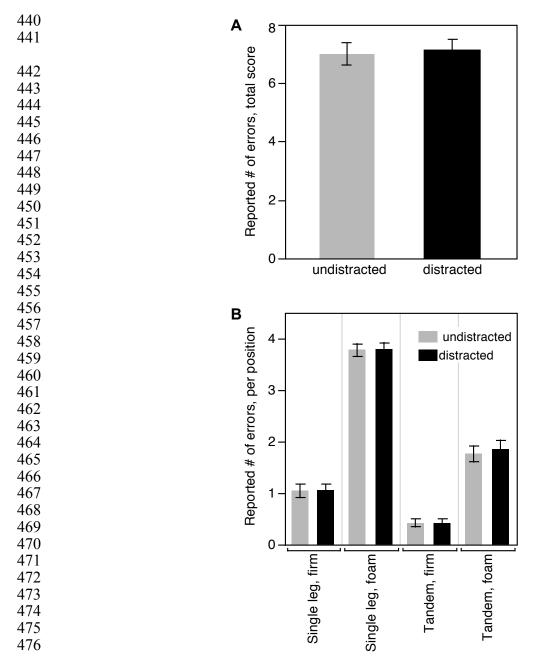
Stance Position	ICC	CI
Double Leg Firm	0.00 ^a	0.000 ^a
Single Leg Firm	0.89	0.811-0.953
Tandem Firm	0.91	0.832-0.959
Double Leg Foam	0.00 ^a	0.000 ^a
Single Leg Foam	0.64	0.465-0.818
Tandem Foam	0.92	0.862-0.968
Total Score	0.92	0.845-0.964
^a Both double leg stance ICCs	and Cle could not be calcule	ated due to a lack of errors

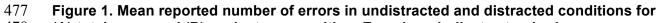
Both double leg stance ICCs and CIs could not be calculated due to a lack of errors.

429

Stance Position	ICC	CI
Double Leg Firm	0.00 ^a	0.000 ^a
Single Leg Firm	0.88	0.753-0.953
Tandem Firm	0.64	0.368-0.837
Double Leg Foam	0.00 ^a	0.000 ^a
Single Leg Foam	0.81	0.594-0.922
Tandem Foam	0.89	0.729-0.959
Total Score	0.89	0.772-0.958
B. Undistracted Inter-Rate	r Reliability ICC's and CI's;	Time 2
Stance Position	ICC	CI
Double Leg Firm	0.00 ^a	0.000 ^a
Single Leg Firm	0.90	0.772-0.961
Tandem Firm	0.92	0.826-0.968
Double Leg Foam	0.00 ^a	0.000 ^a
Single Leg Foam	0.70	0.424-0.874
Tandem Foam	0.76	0.540-0.898
Total Score	0.88	0.658-0.956
C. Undistracted Intra-Rate		
Stance Position	ICC	CI
Double Leg Firm	0.00 ^a	0.000 ^a
Single Leg Firm	0.90	0.805-0.956
Tandem Firm	0.79	0.647-0.907
Double Leg Foam	0.00 ^a	0.000 ^a
Single Leg Foam	0.73	0.545-0.878
Tandem Foam	0.82	0.679-0.921
Total Score	0.89	0.777-0.952
D. Distracted Inter-Rater R		
Stance Position	ICC	CI
Double Leg Firm	0.00 ^a	0.000 ^a
Single Leg Firm	0.79	0.582-0.914
Tandem Firm	0.66	0.392-0.848
Double Leg Foam	0.00 ^a	0.000 ^a
Single Leg Foam	0.43	0.072-0.733
Tandem Foam	0.94	0.858-0.975
Total Score	0.80	0.400-0.930
E. Undistracted vs. Distrac	ted Intra-Rater Reliability IC	
Stance Position	ICC	CI
Double Leg Firm	0.00 ^a	0.000 ^a
Single Leg Firm	0.86	0.743-0.939
Tandem Firm	0.79	0.637-0.904
Double Leg Foam	0.00 ^a	0.000 ^a
Single Leg Foam	0.55	0.294-0.776
Tandem Foam	0.91	0.823-0.960

438 439 a lack of errors. Both double leg stance ICCs and CIs could not be calculated due το





- 478 (A) total score and (B) each stance position. Error bars indicate standard error.